## STATISTICAL SYNTHESIS OF ALGORITHMS FOR THE RECEPTION OF MULTICOMPONENT DISCRETE SIGNALS

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Optimal algorithms have been derived for the joint nonlinear filtering of multicomponent discrete-continuous Markovian processes. Optimal reception algorithms have been synthesized for double phase-shift and frequency phase-shift telegraphy in the presence of a random initial phase of the received signal. The singularities of the resulting block diagrams have been discussed.

Let the received wave have the form

$$r(t) = s(t; \lambda, \alpha) + n(t), \qquad (1)$$

where  $s(t; \lambda, \alpha)$  is a signal depending on a continuous random parameter  $\lambda(t)$  and on a vector of discrete parameters  $\alpha = \{\alpha^{(1)}, \alpha^{(2)}, ..., \alpha^{(M)}\}$ ; n(t) is normal white noise with a zero average and a correlation function  $B(\tau) = (N_0/2)\delta(\tau)$ ;  $N_0$  is the spectral density of the power; the processes  $\lambda$ ,  $\alpha^{(1)}$ ,  $\alpha^{(2)}$ , ...,  $\alpha^{(M)}$  are statistically independent. In the general case, the parameters which are unknown and subject to extraction -  $\lambda(t)$ ,  $\alpha^{(1)}$ ,  $\alpha^{(2)}$ , ...,  $\alpha^{(M)}$  - are related to the signal by nonlinear functional relationships.

Assume that  $\lambda(t)$  is a normal Markovian process, while the components of the vector  $\alpha$  are discrete Markovian processes having n possible states.

Based on the theory of conditional Markovian processes (see [1]), we write the equation for the normalized a posteriori probability density of the parameters  $\lambda$ ,  $\alpha^{(1)}$ ,  $\alpha^{(2)}$ , ...,  $\alpha^{(M)}$ .

$$\frac{\partial}{\partial t} W(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) = L_{pr} W(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) + \\ + \sum_{i=1}^{n} a_{il}^{(1)} W(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) + \sum_{i=1}^{n} a_{ik}^{(2)} W(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) + ...$$
(2)  
$$\dots + \sum_{i=1}^{n} a_{ll}^{(M)} W(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) + [F(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) - \langle F \rangle] \times \\ \times W(t; \lambda, \alpha_{j}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}),$$

where j, k, l = 1, 2, ..., n;  $L_{Pr}$  is the Fokker-Planck-Kolmogorov operator which describes the transformations of the a priori probability density;  $a_{ij}^{(1)}, a_{ik}^{(2)}, ..., a_{il}^{(M)}$  are coefficients which are determined by the local values of the transition probabilities of the corresponding discrete processes  $\alpha^{(1)}, \alpha^{(2)}, ..., \alpha^{(M)}$ ;

$$F(t; \lambda, \alpha_{i}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) = \frac{2}{N_{0}} r(t) s(t; \lambda, \alpha_{i}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)});$$

$$\langle F \rangle = \sum_{j=1}^{n} \sum_{k=1}^{n} ... \sum_{l=1}^{n} \int_{\lambda} F(t; \lambda, \alpha_{i}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) \times \times W(t; \lambda, \alpha_{i}^{(1)}, \alpha_{k}^{(2)}, ..., \alpha_{l}^{(M)}) d\lambda.$$
(3)

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